

# Model parameter estimation with trace forecast likelihood

Ivan Svetunkov and Nikolaos Kourentzes

LCF presentation

7th June 2016



Lancaster University  
Management School

Lancaster Centre for  
Forecasting



LCF



# Motivation

Parameters estimation is a key element of forecasting.

Poor estimation  $\rightarrow$  poor forecasts.

Correct estimation leads to more accurate forecast.

It also decreases the uncertainty.



# Conventional Estimation methods

The conventional estimation methods is based on MSE:

$$\text{MSE} = \frac{1}{T} \sum_{t=1}^T e_{t+1|t}^2 \quad (1)$$

where  $e_{t+1|t} = y_{t+1} - \hat{y}_{t+1}$

MSE – “Mean Squared Error”.



If the errors in the model are distributed normally, than using (1) is equivalent to maximising the following log-likelihood function (Hyndman et al., 2008):

$$\ell(\theta, \hat{\sigma}^2 | \mathbf{Y}) = -\frac{T}{2} (\log(2\pi e) + \log \hat{\sigma}^2) \quad (2)$$

where  $\hat{\sigma}^2$  is the estimated variance of residuals of the model,  $\theta$  is a vector of parameters of the model.

This implies that we look at conditional distribution of one-step-ahead forecast error.



## Advanced estimation methods

Sometimes the forecasting task is aligned to estimation:

$$\text{MSE}_h = \frac{1}{T} \sum_{t=1}^T e_{t+h|t}^2 \quad (3)$$

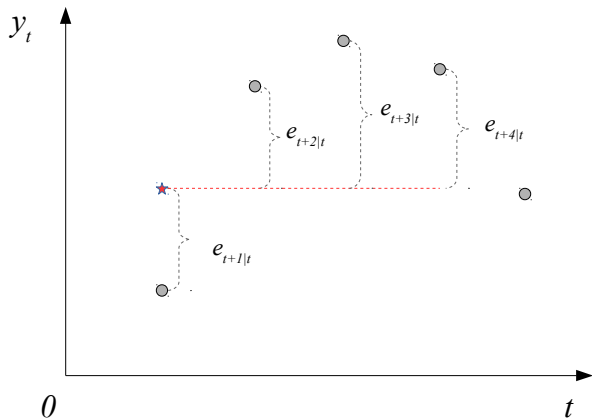
or:

$$\text{MSTFE} = \frac{1}{T} \sum_{t=1}^T \sum_{j=1}^h e_{t+j|t}^2 \quad (4)$$

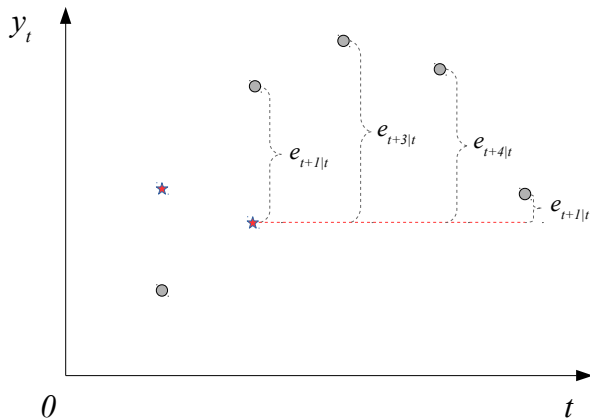
MSTFE – “Mean Squared Trace Forecast Error”.



These cost functions imply that we produce  $h$ -steps ahead forecasts from each observation:



These cost functions imply that we produce  $h$ -steps ahead forecasts from each observation:



$MSE_h$  produces robust estimates of parameters.

(???)

The forecast accuracy increases.

(????)

MSTFE is consistent.

(?)

BUT!

The efficiency of estimates of  $MSE_h$  is low.

(?)

? demonstrate on a set of 170 time series that the forecast accuracy using  $MSE_h$  is lower than using MSE.





## Problems:

- The results are ambiguous;
- Estimates of parameters are inefficient;
- Estimates of parameters could be unstable;
- Nobody has ever explained why  $MSE_h$  and MSTFE work / don't work;
  
- There is no likelihood function for both  $MSE_h$  and MSTFE;
- Model selection using  $MSE_h$  and MSTFE is really tricky (??);



## Why they work

It can be shown that MSE is proportional to variance of one-step-ahead error.

$MSE_h$  is then proportional to variance of h-step-ahead error.

MSTFE is in fact the sum of  $MSE_h$  .

Using state-space approach (Snyder, 1985), variance of h-step-ahead error is:

$$\sigma_h^2 = \sigma_1^2 \left( 1 + \sum_{j=1}^{h-1} c_j^2 \right). \quad (5)$$



This means that minimising  $MSE_h$  (or MSTFE) in general leads to:

1. decrease of variance of one-step-ahead error,
2. shrinkage of values of smoothing parameters towards zero,

Examples:

$$\text{ETS(A,N,N): } c_j = \alpha; \sigma_h^2 = \sigma_1^2 (1 + (h - 1)\alpha^2).$$

$$\text{ETS(A,A,N): } c_j = \alpha + \beta j; \sigma_h^2 = \sigma_1^2 \left(1 + \sum_{j=1}^{h-1} (\alpha + \beta j)^2\right).$$



This is root of the problem and main advantage of  $MSE_h$  and MSTFE.

If model is wrong, shrinkage allows to get rid of redundant parameters.

If model is correct, the parameters “overshrink”.

The shrinkage effect becomes stronger when  $h$  increases.



# Solution – Trace Forecast Likelihood (TFL)

Let's derive likelihood for multistep cost function.  
We need to study multivariate distribution of errors:

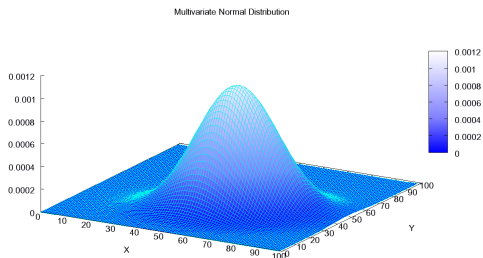


Figure: Multivariate normal distribution.



Based on multivariate normal distribution, we have (skipping derivations):

$$\ell(\theta, \hat{\Sigma} | \mathbf{Y}) = -\frac{T}{2} \left( h \log(2\pi e) + \log |\hat{\Sigma}| \right) \quad (6)$$

Looks similar to:

$$\ell(\theta, \hat{\sigma}_1^2 | \mathbf{Y}) = -\frac{T}{2} \left( \log(2\pi e) + \log \hat{\sigma}_1^2 \right) \quad (7)$$

Model selection can now be done using AIC, AICc, BIC, ...



$\Sigma$  is covariance matrix that has the structure:

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{1,2} & \dots & \sigma_{1,h} \\ \sigma_{1,2} & \sigma_2^2 & \dots & \sigma_{2,h} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{1,h} & \sigma_{2,h} & \dots & \sigma_h^2 \end{pmatrix}, \quad (8)$$

Note that  $MSE_h \propto \sigma_h^2$ , which makes it a special case of  $\Sigma$ .

And MSTFE is just the trace of  $\Sigma$ .



What does min of  $|\Sigma|$  mean?

Example with  $h = 2$ :

$$|\Sigma| = \begin{vmatrix} \sigma_1^2 & \sigma_{1,2} \\ \sigma_{1,2} & \sigma_2^2 \end{vmatrix} = \sigma_1^2 \sigma_2^2 - \sigma_{1,2}^2 \quad (9)$$

Minimising determinant of  $|\Sigma|$  will:

1. decrease variances,
2. increase covariances.





Covariance between  $i$  and  $j$  errors is equal to:

$$\sigma_{i,j} = \sigma_1^2 \left( c_{i,j} + \sum_{l=1}^{i-1} c_{l,j} c_{l,i} \right). \quad (10)$$

So  $\log |\Sigma|$  can be rewritten as a function of variances and parameters:

$$\log |\Sigma| = h \log \sigma_1^2 + \log |\mathbf{C}| \quad (11)$$

$\mathbf{C}$  depends on  $c_j$  only (thus depends on smoothing parameters).

This means that we shrink parameters...

...but shrinkage effect is weakened.

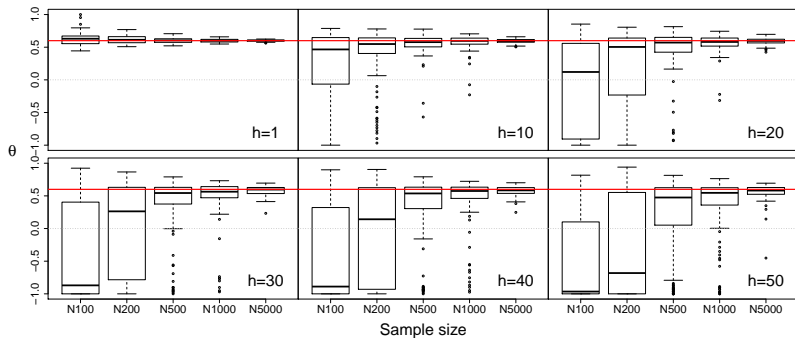


We have conducted a simulation experiment:

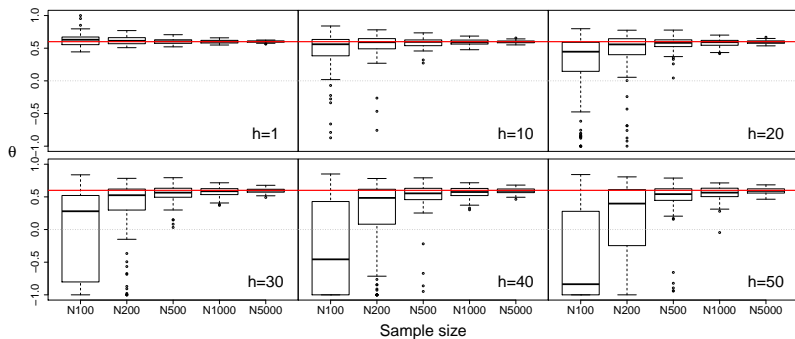
- Generated data using ARIMA(0,1,1).
- Applied correct model and wrong model.
- Applied MSE,  $MSE_h$ , MSTFE and TFL.
- With  $h=\{1, 10, 20, 30, 40, 50\}$ .
- Wrote down the parameters...



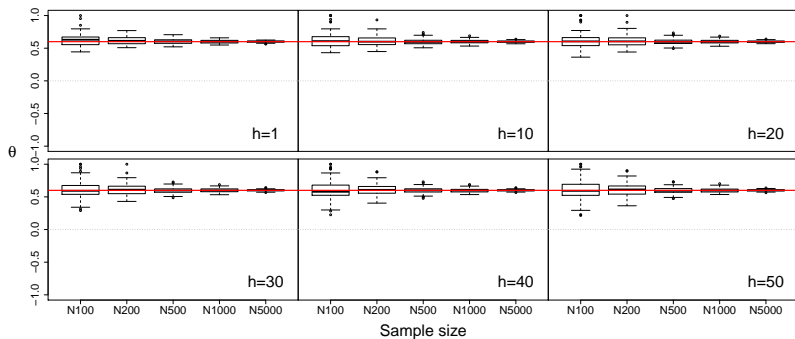
# Simulations. ARIMA(0,1,1). Correct model, $MSE_h$



# Simulations. ARIMA(0,1,1). Correct model, MSTFE



# Simulations. ARIMA(0,1,1). Correct model, TFL



# Conclusions

- Multiple steps ahead objective functions imply shrinkage of parameters;
- Parameters of ETS and ARIMA shrink, parameters of regressions do not;
- This gives robustness to models and help in long-term forecasting;
- Parameters may overshrink when estimated using  $MSE_h$  and MSTFE;



# Conclusions

- Trace Forecast Likelihood (TFL) do not overshrink the parameters;
- TFL gives consistent, efficient and unbiased estimates of parameters;
- Model selection with TFL is possible.

TFL is brilliant in theory!

**How to make it work in practice?...**



# Thank you for your attention!

Ivan Svetunkov

[i.svetunkov@lancaster.ac.uk](mailto:i.svetunkov@lancaster.ac.uk)

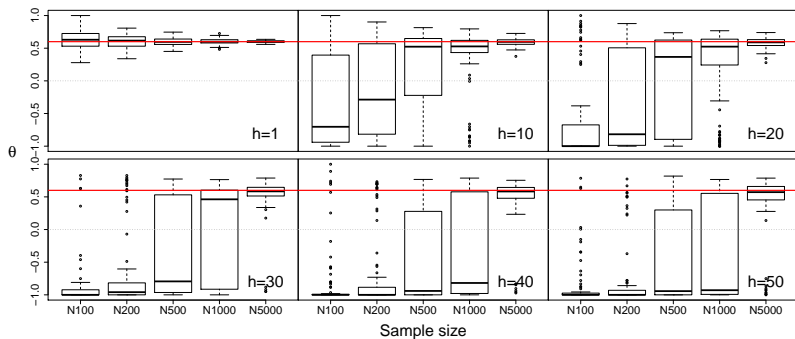


Lancaster University  
Management School

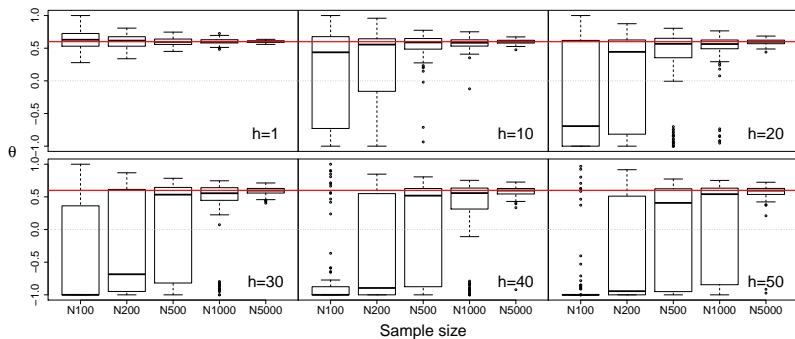
Lancaster Centre for  
Forecasting



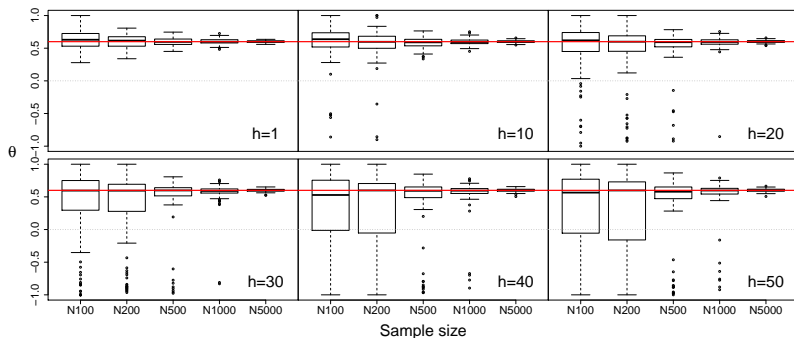
# Simulations. ARIMA(0,1,1). Wrong model, $MSE_h$



# Simulations. ARIMA(0,1,1). Wrong model, MSTFE



# Simulations. ARIMA(0,1,1). Wrong model, TFL



- Bhansali, R., 1996. Asymptotically efficient autoregressive model selection for multistep prediction. *Annals of the Institute of Statistical Mathematics* 48 (3), 577–602.
- Bhansali, R., 1997. Direct autoregressive predictors for multistep prediction: Order selection and performance relative to the plug in predictors. *Statistica Sinica* 7, 425–449.
- Clements, M. P., Hendry, D. F., 2008. *Multi-Step Estimation for Forecasting*.
- Hyndman, R. J., Koehler, A., Ord, K., Snyder, R., 2008. *Forecasting with Exponential Smoothing*. Springer Berlin Heidelberg.
- Marcellino, M., Stock, J. H., Watson, M. W., 2006. A comparison of direct and iterated multistep AR methods for forecasting macroeconomic time series. *Journal of Econometrics* 135 (1-2), 499–526.
- McElroy, T., Wildi, M., 2013. Multi-step-ahead estimation of time



series models. *International Journal of Forecasting* 29 (3), 378–394.

Snyder, R. D., 1985. Recursive Estimation of Dynamic Linear Models. *Journal of the Royal Statistical Society, Series B (Methodological)* 47 (2), 272–276.

Taylor, J. W., 2008. An evaluation of methods for very short-term load forecasting using minute-by-minute British data. *International Journal of Forecasting* 24 (4), 645–658.

Tiao, G., Xu, D., 1993. Robustness of maximum likelihood estimates for multi-step predictions: the exponential smoothing case. *Biometrika* 80 (3), 623–641.

Weiss, A., Andersen, A. P., 1984. Estimating Time Series Models Using the Relevant Forecast Evaluation Criterion. *Journal of the Royal Statistical Society. Series A (General)* 147 (3), 484.

Weiss, A. A., 1991. Multi-step estimation and forecasting in dynamic models. *Journal of Econometrics* 48, 135–149.



Xia, Y., Tong, H., 2011. Feature matching in time series modeling. *Statistical Science* 26 (1), 21–46.

